Noise Enhanced Hypothesis-Testing According to Restricted Neyman-Pearson Criterion

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Abstract

Noise enhanced hypothesis-testing is studied according to the restricted Neyman-Pearson (NP) criterion. First, a problem formulation is presented for obtaining the optimal probability distribution of additive noise in the restricted NP framework. Then, sufficient conditions for improvability and nonimprovability are derived in order to specify if additive noise can or cannot improve detection performance over scenarios in which no additive noise is employed. Also, for the special case of a finite number of possible parameter values under each hypothesis, it is shown that the optimal additive noise can be represented by a discrete random variable with a certain number of point masses. In addition, particular improvability conditions are derived for that special case. Finally, theoretical results are provided for a numerical example and improvements via additive noise are illustrated.

Keywords: Detection, composite hypothesis, noise benefits, stochastic resonance, restricted

Neyman-Pearson.

1. Introduction

Recently, performance improvements obtained via "noise" have been investigated for various problems in the literature ([2] and references therein). Although increasing noise levels or injecting additive noise to a system usually results in degraded performance, it can also lead to performance enhancements in some

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cases. Enhancements obtained via noise can, for instance, be in the form of increased signal-to-noise ratio (SNR), mutual information or detection probability, or in the form of reduced average probability of error [2]-[11].

In hypothesis-testing problems, additive noise can be used to improve performance of a suboptimal detector according to Bayesian, minimax, and Neyman-Pearson (NP) criteria. In [6], the Bayesian criterion is considered under uniform cost assignment, and it is shown that the optimal noise that minimizes the probability of decision error has a constant value. The study in [9] obtains optimal additive noise for suboptimal variable detectors according to the Bayesian and minimax criteria based on the results in [3] and [6]. In [8], noise enhanced *M*-ary composite hypothesis-testing is studied in the presence of partial prior information, and optimal additive noise is investigated according to average and worst-case Bayes risk criteria. In [7], noise enhanced hypothesis-testing is treated in the restricted Bayesian framework, which generalizes the Bayesian and minimax criteria and covers them as special cases [12, 13].

In the NP framework, additive noise can be utilized to increase detection probability of a suboptimal detector under a constraint on false-alarm probability [3, 10, 11, 14]. In [10], an example is provided to illustrate improvements in detection probability due to additive independent noise for the problem of detecting a constant signal in Gaussian mixture noise. A theoretical framework is established in [3] for noise enhanced hypothesis-testing according to the NP criterion, and sufficient conditions are obtained for improvability and nonimprovability of a suboptimal detector via additive noise. In addition, it is shown that optimal additive noise can be realized by a randomization between at most two different signal levels. Noise enhanced detection in the NP framework is studied also in [11], which provides an optimization theoretic framework, and proves the two point mass structure of the optimal additive noise probability distribution.

Noise benefits are studied also for *composite* hypothesis-testing problems, in which there exist multiple possible distributions, hence, multiple parameter values, under each hypothesis [15]. Such problems are encountered in various scenarios such as radar systems, noncoherent communications receivers, and spectrum sensing in cognitive radio networks [15]-[17]. Noise enhanced hypothesis-testing is investigated for composite hypothesis-testing problems according to the Bayesian, NP, and restricted Bayesian criteria in [7, 8, 18]. However, no studies have considered the noise enhanced hypothesis-testing problem according

to the *restricted NP* criterion, which focuses on composite hypothesis-testing problems in the presence of uncertainty in the prior probability distribution under the alternative hypothesis. In the restricted NP framework, the aim is to maximize the average detection probability under constraints on the worst-case detection and false-alarm probabilities [12, 19]. Since prior information may not be perfect in practice, the average detection probability, which is calculated based on the prior distribution under the alternative hypothesis, may not be accurate. Therefore, imposing a constraint on the worst-case detection probability guarantees a minimum detection performance even for the least favorable prior distribution. Hence, the restricted NP approach can have important benefits compared to the NP approach (which aims to maximize the average detection probability under a false-alarm constraint only) when the prior information is not perfect.

In this study, noise enhanced detection is investigated for composite hypothesis-testing problems according to the restricted NP criterion. A formulation is provided for obtaining the probability distribution of the optimal additive noise in the restricted NP framework. Also, sufficient conditions of improvability and nonimprovability are derived in order to determine when the use of additive noise can or cannot improve performance of a given detector according to the restricted NP criterion. In addition, a special case in which there exist finitely many possible values of the unknown parameter under each hypothesis is considered, and the optimal additive noise is shown to correspond to a discrete random variable with a certain number of point masses in that scenario. Furthermore, particular improvability conditions are derived for that special case. Finally, a numerical example is presented to illustrate improvements obtained via additive noise and to provide applications of the improvability conditions. Since a generic composite hypothesis-testing problem with prior distribution uncertainty is investigated in this study, the results can be considered to generalize the previous studies in the literature [3, 11, 18].

The remainder of the manuscript is organized as follows. In Section 2, the noise enhanced hypothesistesting problem is formulated according to the restricted NP criterion, and improvability and nonimprovability conditions are results. In Section 3, the special case with finitely many possible values for the unknown parameter is considered, and particular results are obtained regarding the probability distribution of the optimal additive noise and sufficient conditions for improvability. A numerical example is presented in Section 4 to investigate theoretical results. Finally, concluding remarks are made in Section 5.

2. Noise Enhanced Detection in Restricted NP Framework

We consider a binary composite hypothesis-testing problem formulated as

$$\mathcal{H}_0 : p_{\theta}^{\mathbf{X}}(\mathbf{x}), \ \theta \in \Lambda_0, \qquad \mathcal{H}_1 : p_{\theta}^{\mathbf{X}}(\mathbf{x}), \ \theta \in \Lambda_1$$
(1)

where $p_{\theta}^{\mathbf{X}}(\cdot)$ denotes the probability density function (p.d.f.) of observation \mathbf{X} for a given value of the parameter, $\Theta = \theta$, the observation (measurement), \mathbf{x} , is a *K*-dimensional vector (i.e., $\mathbf{x} \in \mathbb{R}^{K}$), and Λ_{i} is the set of possible parameter values under \mathcal{H}_{i} for i = 0, 1 [15]. Parameter sets Λ_{0} and Λ_{1} are disjoint, and their union forms the parameter space Λ ; that is, $\Lambda = \Lambda_{0} \cup \Lambda_{1}$.

In this study, we consider a practical scenario in which there exists *imperfect* prior information about the parameter. In particular, we assume that the prior probability distribution of the parameter under each hypothesis is known with some *uncertainty* [20]. Let $w_0(\theta)$ and $w_1(\theta)$ represent the *imperfect* prior probability distributions of parameter θ under \mathcal{H}_0 and \mathcal{H}_1 , respectively. These probability distributions may differ from the true prior probability distributions, which are not known by the designer. For instance, $w_0(\theta)$ and $w_1(\theta)$ can be obtained via estimation based on previous decisions (experience). Then, uncertainty is related to estimation errors, and a higher amount of uncertainty is observed as estimation errors increase [19].

For theoretical analysis, we consider a generic decision rule (detector), which is expressed as

$$\phi(\mathbf{x}) = i, \quad \text{if } \mathbf{x} \in \Gamma_i , \tag{2}$$

for i = 0, 1, where Γ_0 and Γ_1 form a partition of the observation space Γ . The aim in this study is to investigate the effects of adding independent "noise" to inputs of given generic detectors as in (2) and to obtain optimal probability distributions of such additive "noise" in the restricted NP framework. As investigated in recent studies such as [2, 3, 7, 9, 10, 11], addition of independent noise to observations can improve detection performance of suboptimal detectors in some cases.

Let **n** denote the "noise" component that is added to original observation **x**. Then, the noise modified observation is formed as $\mathbf{y} = \mathbf{x} + \mathbf{n}$, where **n** has a p.d.f. denoted by $p_{\mathbf{N}}(\cdot)$. The detector in (2) uses the noise

modified observation \mathbf{y} in order to make a decision. As in [3, 7, 11], we assume that the detector in (2) is fixed, and that the only way of enhancing the performance of the detector is to optimize the additive noise component, \mathbf{n} .

According to the restricted NP criterion [12, 19], the optimal additive noise should maximize the average detection probability under constraints on the worst-case detection and false-alarm probabilities. Therefore, the probability distribution of the optimal additive noise can be obtained from the solution of the following optimization problem:

$$\max_{PN(\cdot)} \int_{\Lambda_1} P_D^{\mathbf{y}}(\phi; \theta) w_1(\theta) d\theta$$

subject to $P_D^{\mathbf{y}}(\phi; \theta) \ge \beta, \quad \forall \theta \in \Lambda_1$
 $P_F^{\mathbf{y}}(\phi; \theta) \le \alpha, \quad \forall \theta \in \Lambda_0$ (3)

where $P_D^{\mathbf{y}}(\phi;\theta)$ and $P_F^{\mathbf{y}}(\phi;\theta)$ denote respectively the detection and false-alarm probabilities of a given decision rule ϕ , which employs the noise modified observation \mathbf{y} , for a given value of $\Theta = \theta, \beta$ is the lower limit on the worst-case detection probability, α is the false-alarm constraint, and $w_1(\theta)$ is the imperfect prior distribution of the parameter under hypothesis \mathcal{H}_1 . The objective function in (3) corresponds to the average detection probability based on the imperfect prior distribution; that is, $\int_{\Lambda_1} P_D^{\mathbf{y}}(\phi;\theta) w_1(\theta) d\theta = E\{P_D^{\mathbf{y}}(\phi;\Theta)\} \triangleq P_D^{\mathbf{y}}(\phi)$. In addition, $P_D^{\mathbf{y}}(\phi;\theta)$ and $P_F^{\mathbf{y}}(\phi;\theta)$ can be expressed as

$$P_{\rm D}^{\mathbf{y}}(\phi;\theta) = \mathbb{E}\left\{\phi(\mathbf{Y}) \mid \Theta = \theta\right\} = \int_{\Gamma} \phi(\mathbf{y}) \, p_{\theta}^{\mathbf{Y}}(\mathbf{y}) \, d\mathbf{y} \quad \theta \in \Lambda_1 \tag{4}$$

$$P_{F}^{\mathbf{y}}(\phi;\theta) = \mathbb{E}\left\{\phi(\mathbf{Y}) \mid \Theta = \theta\right\} = \int_{\Gamma} \phi(\mathbf{y}) \, p_{\theta}^{\mathbf{Y}}(\mathbf{y}) \, d\mathbf{y} \quad \theta \in \Lambda_{0}$$
(5)

where $p_{\theta}^{Y}(\cdot)$ is the p.d.f. of the noise modified observation for a given value of $\Theta = \theta$.

In order to express the optimization problem in (3) more explicitly, we first manipulate $P_D^y(\phi; \theta)$ in (4)

as follows:

$$P_{\rm D}^{\mathbf{y}}(\phi;\theta) = \int_{\Gamma} \int_{\mathbb{R}^K} \phi(\mathbf{y}) \, p_{\theta}^{\mathbf{X}}(\mathbf{y} - \mathbf{n}) \, p_{\mathbf{N}}(\mathbf{n}) \, d\mathbf{n} \, d\mathbf{y}$$
(6)

$$= \int_{\mathbb{R}^{K}} p_{\mathbf{N}}(\mathbf{n}) \left[\int_{\Gamma} \phi(\mathbf{y}) p_{\theta}^{\mathbf{X}}(\mathbf{y} - \mathbf{n}) \, d\mathbf{y} \right] d\mathbf{n}$$
(7)

$$\triangleq \int_{\mathbb{R}^K} p_{\mathbf{N}}(\mathbf{n}) F_{\theta}(\mathbf{n}) d\mathbf{n}$$
(8)

$$= \mathbf{E}\{F_{\theta}(\mathbf{N})\}\tag{9}$$

for $\theta \in \Lambda_1$, where the independence of **X** and **N** is used to obtain (6) from (4), and F_{θ} is defined as

$$F_{\theta}(\mathbf{n}) \triangleq \int_{\Gamma} \phi(\mathbf{y}) \, p_{\theta}^{\mathbf{X}}(\mathbf{y} - \mathbf{n}) \, d\mathbf{y}.$$
⁽¹⁰⁾

Note that $F_{\theta}(\mathbf{n})$ corresponds to the detection probability for a given value of $\theta \in \Lambda_1$ and for a constant value of additive noise, $\mathbf{N} = \mathbf{n}$. Therefore, for $\mathbf{n} = \mathbf{0}$, $F_{\theta}(\mathbf{0}) = P_{\mathrm{D}}^{\mathbf{x}}(\phi; \theta)$ is obtained; that is, $F_{\theta}(\mathbf{0})$ is equal to the detection probability of the decision rule for a given value of $\theta \in \Lambda_1$ and for the original observation \mathbf{x} .

Based on similar manipulations as in (6)-(9), $P_F^y(\phi; \theta)$ in (5) can be expressed as

$$\mathbf{P}_{\mathrm{F}}^{\mathbf{y}}(\boldsymbol{\phi};\boldsymbol{\theta}) = \mathrm{E}\{G_{\boldsymbol{\theta}}(\mathbf{N})\}\tag{11}$$

for $\theta \in \Lambda_0$, where

$$G_{\theta}(\mathbf{n}) \triangleq \int_{\Gamma} \phi(\mathbf{y}) \, p_{\theta}^{\mathbf{X}}(\mathbf{y} - \mathbf{n}) \, d\mathbf{y}.$$
(12)

Note that $G_{\theta}(\mathbf{n})$ defines the false alarm probability for a given value of $\theta \in \Lambda_0$ and for a constant value of additive noise, $\mathbf{N} = \mathbf{n}$. Hence, for $\mathbf{n} = \mathbf{0}$, $G_{\theta}(\mathbf{0}) = P_F^{\mathbf{x}}(\phi; \theta)$ is obtained; that is, $G_{\theta}(\mathbf{0})$ is equal to the false alarm probability of the decision rule for a given value of $\theta \in \Lambda_0$ and for the original observation \mathbf{x} .

From (9) and (11), the optimization problem in (3) can be reformulated as

$$\max_{p_{\mathbf{N}}(\cdot)} \int_{\Lambda_{1}} \mathbb{E}\{F_{\theta}(\mathbf{N})\} w_{1}(\theta) d\theta$$

subject to
$$\min_{\theta \in \Lambda_{1}} \mathbb{E}\{F_{\theta}(\mathbf{N})\} \ge \beta$$
$$\max_{\theta \in \Lambda_{0}} \mathbb{E}\{G_{\theta}(\mathbf{N})\} \le \alpha$$
(13)

In addition, based on the following definition,

$$F(\mathbf{n}) \triangleq \int_{\Lambda_1} F_{\theta}(\mathbf{n}) w_1(\theta) \, d\theta \,, \tag{14}$$

the optimization problem in (13) can be expressed in the following simpler form:

$$\max_{p_{\mathbf{N}}(\cdot)} \mathbb{E}\{F(\mathbf{N})\},$$
subject to
$$\min_{\theta \in \Lambda_1} \mathbb{E}\{F_{\theta}(\mathbf{N})\} \ge \beta$$

$$\max_{\theta \in \Lambda_0} \mathbb{E}\{G_{\theta}(\mathbf{N})\} \le \alpha.$$
(15)

Based on the definitions in (10) and (14), it is noted that $F(\mathbf{0}) = P_D^{\mathbf{x}}(\phi)$; that is, $F(\mathbf{0})$ is equal to the average detection probability for the original observation \mathbf{x} (i.e., the average detection probability in the absence of additive noise).

The exact solution of the optimization problem in (15) is very difficult to obtain in general as it requires a search over all possible additive noise p.d.f.s. Hence, an approximate solution can be obtained based on the Parzen window density estimation technique [7, 18, 21]. In particular, the additive noise p.d.f. can be parameterized as

$$p_{\mathbf{N}}(\mathbf{n}) \approx \sum_{l=1}^{L} \mu_l \varphi_l(\mathbf{n})$$
 (16)

where $\mu_l \ge 0$, $\sum_{l=1}^{L} \mu_l = 1$, and $\varphi_l(\cdot)$ is a window function that satisfies $\varphi_l(\mathbf{x}) \ge 0 \ \forall \mathbf{x}$ and $\int \varphi_l(\mathbf{x}) d\mathbf{x} = 1$, for

l = 1, ..., L. A common window function is the Gaussian window, for which $\varphi_l(\mathbf{n})$ is given by the p.d.f. of a Gaussian random vector with a certain mean vector and a covariance matrix. Based on (16), the optimization problem in (15) can be solved over a number of parameters instead of p.d.f.s, which significantly reduces the computational complexity. However, even in that case, the problem is nonconvex in general; hence, global optimization algorithms such as particle swarm optimization (PSO) need to be used [7, 22].

Since the optimization problem in (15) is complex to solve in general, it can be useful to determine *beforehand* if additive noise can or cannot improve the performance of a given detector. For that purpose, we obtain sufficient conditions for which the use of additive noise can or cannot provide any performance improvements compared to the case of not employing any additive noise. To that aim, we first define *improvability* and *nonimprovability* in the restricted NP framework as follows:

Definition 1: According to the restricted NP criterion, a detector is called improvable if there exists additive noise **N** such that $E\{F(\mathbf{N})\} > P_{D}^{\mathbf{x}}(\phi) = F(\mathbf{0})$ and $\min_{\theta \in \Lambda_{1}} P_{D}^{\mathbf{y}}(\phi; \theta) = \min_{\theta \in \Lambda_{1}} E\{F_{\theta}(\mathbf{N})\} \ge \beta$, and $\max_{\theta \in \Lambda_{0}} P_{F}^{\mathbf{y}}(\phi; \theta) = \max_{\theta \in \Lambda_{0}} E\{G_{\theta}(\mathbf{N})\} \le \alpha$. Otherwise, the detector is called nonimprovable.

In other words, for improvability of a detector, there must exist additive noise that increases the average detection probability under the worst-case detection and false-alarm constraints.

According to Definition 1, we first obtain the following nonimprovability condition based on the properties of F_{θ} in (10), G_{θ} in (12), and F in (14).

Proposition 1: Assume that there exits $\theta^* \in \Lambda_0$ ($\theta^* \in \Lambda_1$) such that $G_{\theta^*}(\mathbf{n}) \leq \alpha$ ($F_{\theta^*}(\mathbf{n}) \geq \beta$) implies $F(\mathbf{n}) \leq F(\mathbf{0})$ for all $\mathbf{n} \in S_n$, where S_n is a convex set¹ consisting of all possible values of additive noise \mathbf{n} . If $G_{\theta^*}(\mathbf{n})$ is a convex function ($F_{\theta^*}(\mathbf{n})$ is a concave function), and $F(\mathbf{n})$ is a concave function over S_n , then the detector is nonimprovable.

Proof: The proof is similar to those in [7] and [14]. The convexity of $G_{\theta^*}(\cdot)$ implies that the false alarm probability in (9) is bounded, via Jensen's inequality, as

$$\mathbf{P}_{\mathbf{F}}^{\mathbf{y}}(\boldsymbol{\phi};\boldsymbol{\theta}^*) = \mathbf{E}\{G_{\boldsymbol{\theta}^*}(\mathbf{N})\} \ge G_{\boldsymbol{\theta}^*}\left(\mathbf{E}\{\mathbf{N}\}\right). \tag{17}$$

¹It is reasonable to model S_n as a convex set since convex combination of individual noise components can be obtained via randomization [7, 23].

As $P_F^{\mathbf{y}}(\phi; \theta^*) \leq \alpha$ must hold for improvability, (17) requires that $G_{\theta^*}(E\{\mathbf{N}\}) \leq \alpha$ must be satisfied. Since $E\{\mathbf{N}\} \in S_n, G_{\theta^*}(E\{\mathbf{N}\}) \leq \alpha$ implies that $F(E\{\mathbf{N}\}) \leq F(\mathbf{0})$ due to the assumption in the proposition. Hence,

$$P_{\rm D}^{\mathbf{y}}(\phi) = \mathrm{E}\{F(\mathbf{N})\} \le F\left(\mathrm{E}\{\mathbf{N}\}\right) \le F(\mathbf{0}) , \qquad (18)$$

where the first inequality results from the concavity of *F*. Then, from (17) and (18), it is concluded that whenever the false-alarm constraint is satisfied, the average detection probability can never be higher than that in the absence of additive noise; that is, $P_F^{\mathbf{y}}(\phi; \theta^*) \leq \alpha$ implies $P_D^{\mathbf{y}}(\phi; \theta^*) \leq F(\mathbf{0}) = P_D^{\mathbf{x}}(\phi)$. For this reason, the detector is nonimprovable. Based on similar arguments, the alternative nonimprovability condition in terms of F_{θ} (stated in the parentheses in the proposition) can be proved as well.

The nonimprovability conditions in Proposition 1 can be useful in determining when it is unnecessary to solve the optimization problem in (15). When these conditions are satisfied, additive noise should not be employed in the system at all since it cannot provide any performance improvements according to the restricted NP criterion.

In addition to the nonimprovability conditions in Proposition 1, we obtain sufficient conditions for improvability in the remainder of this section. Assume that $F(\mathbf{x})$, $F_{\theta}(\mathbf{x}) \forall \theta \in \Lambda_1$, and $G_{\theta}(\mathbf{x}) \forall \theta \in \Lambda_0$ are second-order continuously differentiable around $\mathbf{x} = \mathbf{0}$. Then, we define the following functions for notational convenience:

$$g_{\theta}^{(1)}(\mathbf{x}, \mathbf{z}) \triangleq \mathbf{z}^T \nabla G_{\theta}(\mathbf{x}) \tag{19}$$

$$f_{\theta}^{(1)}(\mathbf{x}, \mathbf{z}) \triangleq \mathbf{z}^T \nabla F_{\theta}(\mathbf{x})$$
(20)

$$f^{(1)}(\mathbf{x}, \mathbf{z}) \triangleq \mathbf{z}^T \nabla F(\mathbf{x}) \tag{21}$$

$$g_{\theta}^{(2)}(\mathbf{x}, \mathbf{z}) \triangleq \mathbf{z}^T H(G_{\theta}(\mathbf{x})) \mathbf{z}$$
(22)

$$f_{\theta}^{(2)}(\mathbf{x}, \mathbf{z}) \triangleq \mathbf{z}^T H(F_{\theta}(\mathbf{x})) \mathbf{z}$$
(23)

$$f^{(2)}(\mathbf{x}, \mathbf{z}) \triangleq \mathbf{z}^T H(F(\mathbf{x})) \mathbf{z}$$
(24)

where z is a K-dimensional column vector, and ∇ and H represent the gradient and Hessian operators,

respectively. For example, $\nabla G_{\theta}(\mathbf{x})$ is a *K*-dimensional column vector with its *i*th element being equal to $\frac{\partial G_{\theta}(\mathbf{x})}{\partial x_i}$, where x_i denotes the *i*th component of \mathbf{x} , and $H(G_{\theta}(\mathbf{x}))$ is a $K \times K$ matrix with its element in row *l* and column *i* being given by $\frac{\partial^2 G_{\theta}(\mathbf{x})}{\partial x_i \partial x_i}$.

Based on the preceding definitions, the following proposition provides sufficient conditions for improvability.

Proposition 2: Let \mathcal{L}_0 and \mathcal{L}_1 denote the sets of θ values that maximize $G_{\theta}(\mathbf{0})$ and minimize $F_{\theta}(\mathbf{0})$, respectively. Then the detector is improvable if there exists a K-dimensional vector \mathbf{z} such that one of the following conditions is satisfied for all $\theta_0 \in \mathcal{L}_0$ and $\theta_1 \in \mathcal{L}_1$:

- $f^{(1)}(\mathbf{x}, \mathbf{z}) > 0$, $f^{(1)}_{\theta_1}(\mathbf{x}, \mathbf{z}) > 0$, and $g^{(1)}_{\theta_0}(\mathbf{x}, \mathbf{z}) < 0$ at $\mathbf{x} = \mathbf{0}$.
- $f^{(1)}(\mathbf{x}, \mathbf{z}) < 0$, $f^{(1)}_{\theta_1}(\mathbf{x}, \mathbf{z}) < 0$, and $g^{(1)}_{\theta_0}(\mathbf{x}, \mathbf{z}) > 0$ at $\mathbf{x} = \mathbf{0}$.
- $f^{(2)}(\mathbf{x}, \mathbf{z}) > 0$, $f^{(2)}_{\theta_1}(\mathbf{x}, \mathbf{z}) > 0$, and $g^{(2)}_{\theta_0}(\mathbf{x}, \mathbf{z}) < 0$ at $\mathbf{x} = \mathbf{0}$.

Proof: Please see Appendix A.1.

Proposition 2 implies that under the stated conditions, one can always find a noise p.d.f. that increases the average detection probability under the constraints on the worst case detection and false alarm probabilities. In other words, the conditions in the proposition guarantee the existence of additive noise that improves the detection performance according to the restricted NP criterion.

In addition to the improvability conditions in Proposition 2, we can obtain alternative sufficient conditions for improvability based on the approaches in [3, 7]. For that purpose, we first define two new functions J(t) and H(t) as follows:

$$J(t) \triangleq \sup\left\{F(\mathbf{n}) \mid \max_{\theta \in \Lambda_0} G_{\theta}(\mathbf{n}) = t , \ \mathbf{n} \in \mathbb{R}^K\right\}$$
(25)

$$H(t) \triangleq \inf \left\{ \min_{\theta \in \Lambda_1} F_{\theta}(\mathbf{n}) \mid \max_{\theta \in \Lambda_0} G_{\theta}(\mathbf{n}) = t , \ \mathbf{n} \in \mathbb{R}^K \right\}$$
(26)

which represent, respectively, the maximum average detection probability and the minimum worst-case detection probability for a given value of the maximum false-alarm probability considering constant values of additive noise. As an initial observation from (25) and (26), one can conclude that if there exists $t_0 \leq \alpha$ such that $J(t_0) > F(\mathbf{0})$ and $H(t_0) \geq \beta$, then the detector is improvable, since under such a condition

there exists a noise component \mathbf{n}_0 that satisfies $F(\mathbf{n}_0) > F(\mathbf{0})$, $\min_{\theta \in \Lambda_1} F_{\theta}(\mathbf{n}_0) \ge \beta$ and $\max_{\theta \in \Lambda_0} G_{\theta}(\mathbf{n}_0) \le \alpha$ (i.e., performance improvement can be achieved by adding a constant noise component \mathbf{n}_0 to the observation).

Since improvability of a detector via constant noise component is not very common in practice, the following improvability condition is presented for more practical scenarios.

Proposition 3: Define the minimum value of the detection probability and the maximum value of the false alarm probability in the absence of additive noise as $\tilde{\beta} \triangleq \min_{\theta \in \Lambda_1} P_D^{\mathbf{x}}(\phi; \theta)$ and $\tilde{\alpha} \triangleq \max_{\theta \in \Lambda_0} P_F^{\mathbf{x}}(\phi; \theta)$, respectively, where $\tilde{\beta} \ge \beta$ and $\tilde{\alpha} \le \alpha$. Assume that $H(\tilde{\alpha}) = \tilde{\beta}$, where H is as defined in (26). Then the detector is improvable if J(t) in (25) and H(t) in (26) are second-order continuously differentiable around $t = \tilde{\alpha}$, and satisfy $J''(\tilde{\alpha}) > 0$ and $H''(\tilde{\alpha}) \ge 0$.

Proof: Please see Appendix A.2.

Proposition 3 can be employed in a similar manner to Proposition 2 in order to determine if a given detector is improvable according to the restricted NP framework. The main advantage of Proposition 3 is that J(t) and H(t) are always single-variable functions irrespective of the dimension of the observation vector, which facilitates simple evaluation of the conditions in the proposition. However, in some cases, it can be challenging to obtain an expression for J(t) in (25) and H(t) in (26). On the other hand, Proposition 2 deals directly with $G_{\theta}(\cdot)$, $F_{\theta}(\cdot)$, and $F(\cdot)$ without defining auxiliary functions as in Proposition 3; hence, it can be employed more efficiently in some cases. However, it should also be noted that the functions in Proposition 2 are always *K*-dimensional, which can make the evaluation of the conditions more complex than those in Proposition 3 in some other cases.

3. Special Case: Finitely Many Possible Values for the Parameter

The results obtained in the previous section are generic in the sense that there are no specific restrictions on the parameter sets Λ_0 and Λ_1 corresponding to hypotheses \mathcal{H}_0 and \mathcal{H}_1 , respectively. In this section, we provide more detailed theoretical analysis for the special case in which the parameter sets consist of finitely many elements. Let $\Lambda_0 = \{\theta_{01}, \theta_{02}, \dots, \theta_{0M}\}$ and $\Lambda_1 = \{\theta_{11}, \theta_{12}, \dots, \theta_{1N}\}$.

The most important simplification in this case is that the optimal probability distribution of additive noise can be represented by a discrete probability distribution with at most M + N point masses under mild

conditions as specified in the following proposition.

Proposition 4: Suppose that each component of additive noise is upper and lower bounded by two finite values; that is, $n_j \in [a_j, b_j]$ for j = 1, ..., K where a_j and b_j are finite.² If $F_{\theta_{1i}}(\cdot)$ and $G_{\theta_{0i}}(\cdot)$ are continuous functions, then the p.d.f. of an optimal additive noise can be expressed as

$$p_{\mathbf{N}}(\mathbf{n}) = \sum_{l=1}^{M+N} \lambda_l \,\delta(\mathbf{n} - \mathbf{n}_l) \,, \tag{27}$$

where $\sum_{l=1}^{M+N} \lambda_l = 1$ and $\lambda_l \ge 0$ for $l = 1, 2, \dots, M + N$.

Proof: The proof is omitted since it can be obtained similarly to the proofs of Theorem 4 in [7], Theorem 8 in [18], and Theorem 3 in [3].

Based on Proposition 4, the optimization problem in (15) can be expressed as

$$\max_{\{\lambda_l, \mathbf{n}_l\}_{l=1}^{M+N}} \sum_{l=1}^{M+N} \lambda_l F(\mathbf{n}_l)$$

subject to
$$\min_{\theta \in \Lambda_1} \sum_{l=1}^{M+N} \lambda_l F_{\theta}(\mathbf{n}_l) \ge \beta$$
$$\max_{\theta \in \Lambda_0} \sum_{l=1}^{M+N} \lambda_l G_{\theta}(\mathbf{n}_l) \le \alpha$$
$$\sum_{l=1}^{M+N} \lambda_l = 1, \ \lambda_l \ge 0 \text{ for } l = 1, 2, \dots, M+N$$
(28)

Compared to (15), the optimization problem in (28) has much lower computational complexity in general since it requires optimization over a number of variables instead of over all possible p.d.f.s. However, depending on the number of possible parameter values, M + N, the computational complexity can still be high in some cases.

Next, we obtain sufficient conditions for improvability according to the restricted NP criterion. Let S_{β} (S_{α}) denote the set of indices for which $F_{\theta_{1i}}(\mathbf{0})$ ($G_{\theta_{0i}}(\mathbf{0})$) achieves the minimum value of β (maximum value

²This is a reasonable assumption because additive noise cannot take infinitely large values in practice.

of α), and let $\bar{S}_{\beta}(\bar{S}_{\alpha})$ represent the set of indices with $F_{\theta_{1i}}(\mathbf{0}) > \beta$ ($G_{\theta_{0i}}(\mathbf{0}) < \alpha$); that is,

$$S_{\beta} = \{ i \in \{1, 2, \dots, N\} \mid F_{\theta_{1i}}(\mathbf{0}) = \beta \}$$
(29)

$$\bar{S}_{\beta} = \{ i \in \{1, 2, \dots, N\} \mid F_{\theta_{1}i}(\mathbf{0}) > \beta \}$$
(30)

$$\mathcal{S}_{\alpha} = \{i \in \{1, 2, \dots, M\} \mid G_{\theta_{0i}}(\mathbf{0}) = \alpha\}$$

$$(31)$$

$$\bar{S}_{\alpha} = \{ i \in \{1, 2, \dots, M\} \mid G_{\theta_{0i}}(\mathbf{0}) < \alpha \}.$$
(32)

Note that $S_{\beta} \cup \bar{S}_{\beta} = \{1, 2, ..., N\}$ $(S_{\alpha} \cup \bar{S}_{\alpha} = \{1, 2, ..., M\})$; hence, $F_{\theta_{1i}}(\mathbf{0}) = P_{D}^{\mathbf{x}}(\phi; \theta_{1i}) \ge \beta$ for i = 1, 2, ..., N $(G_{\theta_{0i}}(\mathbf{0}) = P_{F}^{\mathbf{x}}(\phi; \theta_{0i}) \le \alpha$ for i = 1, 2, ..., M.

Based on the functions in (19)-(24), we define new functions as $f_i^{(n)}(\mathbf{x}, \mathbf{z}) \triangleq f_{\theta_{1i}}^{(n)}(\mathbf{x}, \mathbf{z})$ and $g_i^{(n)}(\mathbf{x}, \mathbf{z}) \triangleq g_{\theta_{1i}}^{(n)}(\mathbf{x}, \mathbf{z})$. Also let \mathcal{F}_n and \mathcal{G}_n (n = 1, 2) represent the sets that consist of $f^{(n)}(\mathbf{x}, \mathbf{z})$, $f_i^{(n)}(\mathbf{x}, \mathbf{z})$ for $i \in S_\beta$, and $g_i^{(n)}(\mathbf{x}, \mathbf{z})$ for $i \in S_\beta$; namely,

$$\mathcal{F}_{n} = \left\{ f^{(n)}(\mathbf{x}, \mathbf{z}), f_{i}^{(n)}(\mathbf{x}, \mathbf{z}) \text{ for } i \in \mathcal{S}_{\beta} \right\}$$
(33)

$$\mathcal{G}_n = \left\{ g_i^{(n)}(\mathbf{x}, \mathbf{z}) \text{ for } i \in \mathcal{S}_\alpha \right\},\tag{34}$$

for n = 1, 2. Note that $\mathcal{F}_n(\mathcal{G}_n)$ has $|\mathcal{S}_\beta| + 1$ ($|\mathcal{S}_\alpha|$) elements, where $|\mathcal{S}_\beta|$ ($|\mathcal{S}_\alpha|$) denotes the number of elements in $\mathcal{S}_\beta(\mathcal{S}_\alpha)$. Representing by $\mathcal{F}_n(j)$ ($\mathcal{G}_n(j)$) the *j*th element of $\mathcal{F}_n(\mathcal{G}_n)$, it is noted that $\mathcal{F}_n(1) = f^{(n)}(\mathbf{x}, \mathbf{z})$ and $\mathcal{F}_n(j) = f^{(n)}_{\mathcal{S}_\beta(j-1)}(\mathbf{x}, \mathbf{z})$ for $j = 2, ..., |\mathcal{S}_\beta| + 1$ ($\mathcal{G}_n(j) = g^{(n)}_{\mathcal{S}_\alpha(j)}(\mathbf{x}, \mathbf{z})$ for $j = 2, ..., |\mathcal{S}_\alpha|$), where $\mathcal{S}_\beta(j-1)$ is the (*j* – 1)th element of $\mathcal{S}_\beta(\mathcal{S}_\alpha(j)$ is the *j*th element of \mathcal{S}_α). Furthermore, the following sets are defined for the indices $j \in S_{\beta}$ $(j \in S_{\alpha})$ for which $\mathcal{F}_1(j)$ $(\mathcal{G}_1(j))$ is zero, negative or positive:

$$S_{\beta}^{z} = \left\{ j \in \{1_{\beta}, 2_{\beta}, \dots, (|S_{\beta}| + 1)_{\beta}\} \mid \mathcal{F}_{1}(j) = 0 \right\}$$
(35)

$$S_{\beta}^{n} = \left\{ j \in \{1_{\beta}, 2_{\beta}, \dots, (|S_{\beta}| + 1)_{\beta}\} \mid \mathcal{F}_{1}(j) < 0 \right\}$$
(36)

$$S_{\beta}^{p} = \left\{ j \in \{1_{\beta}, 2_{\beta}, \dots, (|S_{\beta}| + 1)_{\beta}\} \mid \mathcal{F}_{1}(j) > 0 \right\}$$
(37)

$$S_{\alpha}^{z} = \{ j \in \{1_{\alpha}, 2_{\alpha}, \dots, (|S_{\alpha}|)_{\alpha} \} \mid \mathcal{G}_{1}(j) = 0 \}$$
(38)

$$\mathcal{S}_{\alpha}^{n} = \{ j \in \{1_{\alpha}, 2_{\alpha}, \dots, (|\mathcal{S}_{\alpha}|)_{\alpha}\} \mid \mathcal{G}_{1}(j) < 0 \}$$

$$(39)$$

$$\mathcal{S}^{p}_{\alpha} = \{ j \in \{1_{\alpha}, 2_{\alpha}, \dots, (|\mathcal{S}_{\alpha}|)_{\alpha}\} \mid \mathcal{G}_{1}(j) > 0 \}$$

$$\tag{40}$$

where we denote j as j_{α} (j_{β}) in order to emphasize that j is coming from set S_{α} (is not coming from set S_{α}).

In the following proposition, an indicator function $I_A(x)$ is used, which is defined as $I_A(x) = 1$ if $x \in A$ and $I_A(x) = 0$ otherwise. Based on the definitions in (29)-(40), the following proposition provides sufficient conditions for improvability in the restricted NP framework.

Proposition 5: When Λ consists of a finite number of elements, a detector is improvable according to the restricted NP criterion if there exists a K-dimensional vector \mathbf{z} such that the following two conditions are satisfied at $\mathbf{x} = \mathbf{0}$:

- 1. $\mathcal{F}_2(j) > 0$, $\forall j \in S^z_\beta$ and $\mathcal{G}_2(j) < 0$, $\forall j \in S^z_\alpha$.
- 2. One of the following is satisfied:
 - Any three of $|S_{\beta}^{n}|$, $|S_{\beta}^{p}|$, $|S_{\alpha}^{n}|$ and $|S_{\alpha}^{p}|$ is zero, or $|S_{\beta}^{n}| + |S_{\alpha}^{p}| = 0$, or $|S_{\alpha}^{n}| + |S_{\beta}^{p}| = 0$.
 - $|S_{\beta}^{n}| + |S_{\alpha}^{n}|$ is an odd number, $|S_{\beta}^{n}| + |S_{\alpha}^{p}| > 0$, $|S_{\alpha}^{n}| + |S_{\beta}^{p}| > 0$ and

$$\begin{split} & \min_{j \in S_{\beta}^{n} \cup S_{\alpha}^{p}} \left(\mathcal{F}_{2}(j) \mathcal{I}_{S_{\beta}^{n}}(j) + \mathcal{G}_{2}(j) \mathcal{I}_{S_{\alpha}^{p}}(j) \right) \prod_{l \in S_{\beta}^{n} \cup S_{\beta}^{p} \cup S_{\alpha}^{n} \cup S_{\alpha}^{p} \setminus \{j\}} \left(\mathcal{F}_{1}(l) \mathcal{I}_{S_{\beta}^{n} \cup S_{\beta}^{p}}(l) + \mathcal{G}_{1}(l) \mathcal{I}_{S_{\alpha}^{n} \cup S_{\alpha}^{p}}(l) \right) \\ & > \max_{j \in S_{\beta}^{p} \cup S_{\alpha}^{n}} \left(\mathcal{F}_{2}(j) \mathcal{I}_{S_{\beta}^{p}}(j) + \mathcal{G}_{2}(j) \mathcal{I}_{S_{\alpha}^{n}}(j) \right) \prod_{l \in S_{\beta}^{n} \cup S_{\alpha}^{p} \cup S_{\alpha}^{n} \cup S_{\alpha}^{p} \setminus \{j\}} \left(\mathcal{F}_{1}(l) \mathcal{I}_{S_{\beta}^{n} \cup S_{\beta}^{p}}(l) + \mathcal{G}_{1}(l) \mathcal{I}_{S_{\alpha}^{n} \cup S_{\alpha}^{p}}(l) \right). \end{split}$$

(41)

• $|S_{\beta}^{n}| + |S_{\alpha}^{n}|$ is an even number, $|S_{\beta}^{n}| + |S_{\alpha}^{p}| > 0$, $|S_{\alpha}^{n}| + |S_{\beta}^{p}| > 0$ and

$$\min_{j\in\mathcal{S}_{\beta}^{p}\cup\mathcal{S}_{\alpha}^{n}} \left(\mathcal{F}_{2}(j)\mathcal{I}_{\mathcal{S}_{\beta}^{p}}(j) + \mathcal{G}_{2}(j)\mathcal{I}_{\mathcal{S}_{\alpha}^{n}}(j)\right) \prod_{l\in\mathcal{S}_{\beta}^{n}\cup\mathcal{S}_{\beta}^{p}\cup\mathcal{S}_{\alpha}^{n}\cup\mathcal{S}_{\alpha}^{p}\setminus\{j\}} \left(\mathcal{F}_{1}(l)\mathcal{I}_{\mathcal{S}_{\beta}^{n}\cup\mathcal{S}_{\beta}^{p}}(l) + \mathcal{G}_{1}(l)\mathcal{I}_{\mathcal{S}_{\alpha}^{n}\cup\mathcal{S}_{\alpha}^{p}}(l)\right) \\
> \max_{j\in\mathcal{S}_{\beta}^{n}\cup\mathcal{S}_{\alpha}^{p}} \left(\mathcal{F}_{2}(j)\mathcal{I}_{\mathcal{S}_{\beta}^{n}}(j) + \mathcal{G}_{2}(j)\mathcal{I}_{\mathcal{S}_{\alpha}^{p}}(j)\right) \prod_{l\in\mathcal{S}_{\beta}^{n}\cup\mathcal{S}_{\beta}^{p}\cup\mathcal{S}_{\alpha}^{n}\cup\mathcal{S}_{\alpha}^{p}\setminus\{j\}} \left(\mathcal{F}_{1}(l)\mathcal{I}_{\mathcal{S}_{\beta}^{n}\cup\mathcal{S}_{\beta}^{p}}(l) + \mathcal{G}_{1}(l)\mathcal{I}_{\mathcal{S}_{\alpha}^{n}\cup\mathcal{S}_{\alpha}^{p}}(l)\right).$$

$$(42)$$

Proof: Please see Appendix A.3.

According to Proposition 5, whenever the two conditions in the proposition are satisfied, it is guaranteed that the detection performance can be improved via additive noise. Although the expression in the proposition can seem complicated at first, it is noted that, after defining the sets in (29)-(40), it is simple to check the conditions stated in the proposition. An example application of Proposition 5 is provided in the next section.

The following improvability condition can be obtained as a corollary of Proposition 5.

Corollary 1: Assume that $F(\mathbf{x})$, $F_{\theta_{1i}}(\mathbf{x})$, i = 1, 2, ..., N, and $G_{\theta_{0i}}(\mathbf{x})$, i = 1, 2, ..., M are second-order continuously differentiable around $\mathbf{x} = \mathbf{0}$ and that $\min_{i \in \{1, 2, ..., N\}} F_{\theta_{1i}}(\mathbf{0}) > \beta$ and $\max_{i \in \{1, 2, ..., M\}} G_{\theta_{0i}}(\mathbf{0}) < \alpha$. Let \mathbf{f} denote the gradient of $F(\mathbf{x})$ at $\mathbf{x} = \mathbf{0}$. Then, the detector is improvable

- if $\mathbf{f} \neq \mathbf{0}$; or,
- *if* $F(\mathbf{x})$ *is not concave around* $\mathbf{x} = \mathbf{0}$.

Proof: Please see Appendix A.4.

4. Numerical Results

In this section, the binary hypothesis-testing problem considered in [19] is studied in order to illustrate theoretical results in the previous sections. The hypotheses are specified as follows:

$$\mathcal{H}_0 : X = V , \quad \mathcal{H}_1 : X = \Theta + V \tag{43}$$

where $X \in \mathbb{R}$, Θ is the unknown parameter, and *V* is symmetric Gaussian mixture noise that has the following p.d.f.

$$p_V(v) = \sum_{i=1}^{N_m} \omega_i \,\psi_i(v - m_i) \,, \tag{44}$$

where $\omega_i \ge 0$ for $i = 1, ..., N_m$, $\sum_{i=1}^{N_m} \omega_i = 1$, and $\psi_i(x) = 1/(\sqrt{2\pi} \sigma_i) \exp\left(-x^2/(2\sigma_i^2)\right)$ for $i = 1, ..., N_m$. Since noise V is symmetric, its parameters satisfy $m_l = -m_{N_m-l+1}$, $\omega_l = \omega_{N_m-l+1}$ and $\sigma_l = \sigma_{N_m-l+1}$ for $l = 1, ..., \lfloor N_m/2 \rfloor$, where $\lfloor y \rfloor$ denotes the largest integer smaller than or equal to y. (If N_m is an odd number, $m_{(N_m+1)/2}$ is set to zero for symmetry.)

The unknown parameter Θ in (43) is modeled as a random variable with the following p.d.f.

$$w_1(\theta) = \rho \,\delta(\theta - A) + (1 - \rho) \,\delta(\theta + A) \tag{45}$$

where A is a positive constant that is known exactly, whereas ρ is known with some uncertainty. (Please see [19] for the motivations of this model.)

Based on the preceding problem formulation, the parameter sets under \mathcal{H}_0 and \mathcal{H}_1 are specified as $\Lambda_0 = \{0\}$ and $\Lambda_1 = \{-A, A\}$, respectively. Also, the conditional p.d.f. of the original observation X for a given value of $\Theta = \theta$ is obtained as

$$p_{\theta}^{\mathbf{X}}(x) = \sum_{i=1}^{N_m} \frac{\omega_i}{\sqrt{2\pi}\,\sigma_i} \,\exp\!\left(\frac{-(x-\theta-m_i)^2}{2\,\sigma_i^2}\right). \tag{46}$$

Suppose that the following detector is employed.

$$\phi(y) = \begin{cases} 0, & A/2 > y > -A/2 \\ 1, & \text{otherwise} \end{cases},$$
(47)

where y = x + n, with *n* representing the additive noise term. This is a reasonable detector for the model in (43) since noise *V* is zero mean, and Θ is either *A* of -A. Although it is not the optimal detector for the specified problem, it can be employed in practical scenarios due to its simplicity.

From (10), (12), and (14), $F_{\theta_{1i}}$ for $\theta_{11} = A$ and $\theta_{12} = -A$, $G_{\theta_{0i}}$ for $\theta_{01} = 0$, and F can be calculated as follows:

$$F_{A}(n) = \sum_{i=1}^{N_{m}} w_{i} \left(\mathcal{Q} \left(\frac{-A/2 - m_{i} - n}{\sigma_{i}} \right) + \mathcal{Q} \left(\frac{3A/2 + m_{i} + n}{\sigma_{i}} \right) \right),$$
(48)

$$F_{-A}(n) = \sum_{i=1}^{N_{m}} w_{i} \left(\mathcal{Q} \left(\frac{3A/2 - m_{i} - n}{\sigma_{i}} \right) + \mathcal{Q} \left(\frac{-A/2 + m_{i} + n}{\sigma_{i}} \right) \right),$$
(48)

$$G_{0}(n) = \sum_{i=1}^{N_{m}} w_{i} \left(\mathcal{Q} \left(\frac{A/2 - m_{i} - n}{\sigma_{i}} \right) + \mathcal{Q} \left(\frac{A/2 + m_{i} + n}{\sigma_{i}} \right) \right),$$
(48)

where $Q(x) = (1/\sqrt{2\pi}) \int_x^\infty e^{-t^2/2} dt$ is the *Q*-function.

In the numerical example, $N_m = 4$ is considered for the symmetric Gaussian mixture noise, and the mean values of the Gaussian components in the mixture noise are specified as [0.01 0.6 -0.6 -0.01] with the corresponding weights of [0.25 0.25 0.25 0.25]. Also, the variances of the Gaussian components in the mixture noise are assumed to be the same; i.e., $\sigma_i = \sigma$ for $i = 1, ..., N_m$.

In Figures 1, 2, and 3, average detection probabilities are plotted with respect to σ for various values of β in the cases of $\alpha = 0.35$, $\alpha = 0.4$, and $\alpha = 0.45$, respectively, where A = 1 and $\rho = 0.8$. It is observed that the use of additive noise enhances the average detection probability, and significant improvements can be achieved via additive noise for low values of the standard deviation, σ . As the standard deviation increases, the amount of improvement in the average detection probability reduces. In fact, after some values of σ , the constraints on the minimum detection probability or the false alarm probability are not satisfied; hence, the restricted NP solution does not exist after certain values of σ . (Therefore, the curves are plotted up to those specific values in the figures.) Another observation from the figures is that the average detection probabilities decrease as β increases. This is expected since a larger value of β imposes a more strict constraint on the worst-case detection probability (see (3)), which in turn reduces the average detection probability. In other words, there is a tradeoff between β and the average detection probability, which is an essential characteristics of the restricted NP approach [19].

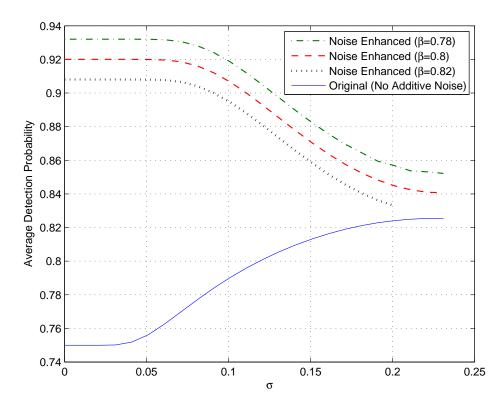


Figure 1: Average detection probability versus σ for various values of β , where $\alpha = 0.35$, A = 1 and $\rho = 0.8$.

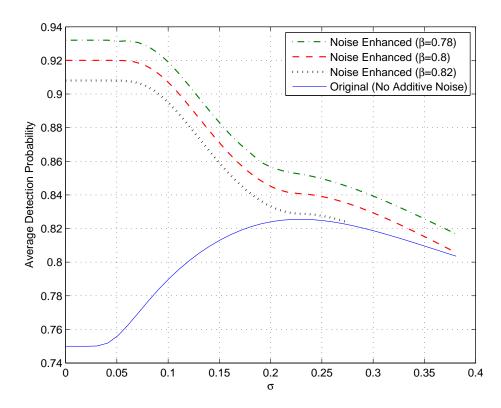


Figure 2: Average detection probability versus σ for various values of β , where $\alpha = 0.4$, A = 1 and $\rho = 0.8$.

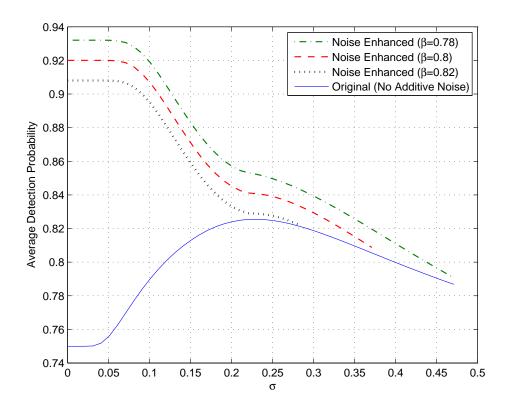


Figure 3: Average detection probability versus σ for various values of β , where $\alpha = 0.45$, A = 1 and $\rho = 0.8$.

Table 1: Optimal additive noise p.d.f.s, in the form of $p_N(n) = \lambda_1 \,\delta(n-n_1) + \lambda_2 \,\delta(n-n_2) + (1-\lambda_1-\lambda_2) \,\delta(n-n_3)$, for various values of σ , where $\beta = 0.82$, $\alpha = 0.35$, A = 1 and $\rho = 0.8$.

σ	λ_1	λ_2	n_1	n_2	<i>n</i> ₃
0	0.4181	0.3019	0.1136	0.4887	-0.4807
0.01	0.5043	0.2157	0.4146	0.1718	-0.4115
0.1	0.6886	0.3114	0.2818	-0.2818	_
0.15	0.6032	0.3968	0.2544	-0.2544	_
0.2	0.5481	0.4519	0.1796	-0.1796	_

Table 2: Optimal additive noise p.d.f.s, in the form of $p_N(n) = \lambda_1 \,\delta(n-n_1) + \lambda_2 \,\delta(n-n_2) + (1-\lambda_1-\lambda_2) \,\delta(n-n_3)$, for various values of σ , where $\beta = 0.8$, $\alpha = 0.4$, A = 1 and $\rho = 0.8$.

σ	λ_1	λ_2	n_1	n_2	<i>n</i> ₃
0	0.6098	0.1902	0.4750	0.2088	-0.2804
0.05	0.5375	0.2624	0.3002	0.2956	-0.2755
0.1	0.7689	0.2311	0.2821	-0.2821	_
0.2	0.6653	0.3347	0.1796	-0.1796	_
0.3	1	_	0.0384	_	—

Tables 1, 2, and 3 illustrate the optimal additive noise p.d.f.s for various values of σ in the cases of $\beta = 0.82$ with $\alpha = 0.35$, $\beta = 0.80$ with $\alpha = 0.40$, and $\beta = 0.78$ with $\alpha = 0.45$ respectively, where A = 1 and $\rho = 0.8$. From Proposition 4, it is known that the optimal additive noise in this example can be represented by a discrete probability distribution with at most three point masses (since $\Lambda_0 = \{0\}$ and $\Lambda_1 = \{-A, A\}$; i.e., M = 1 and N = 2). Therefore, it can be expressed as $p_N(n) = \lambda_1 \delta(n-n_1) + \lambda_2 \delta(n-n_2) + (1-\lambda_1-\lambda_2) \delta(n-n_3)$. It is observed from the tables that the optimal additive noise p.d.f.s have three point masses for certain values of σ , whereas they have two point masses or a single point mass for other σ 's. These results are in accordance with Proposition 4, which states that an optimal p.d.f. can be represented by a probability distribution with at most three point masses for certain.

In order to determine if any of the conditions in Proposition 2 are satisfied for the example above, the numerical values of $f^{(2)}$, $f^{(2)}_{\theta_1}$, and $g^{(2)}_{\theta_0}$ are calculated and tabulated in Table 4.³ It is observed that, in this specific example, $F_{\theta_1}(\mathbf{0})$ has two minimizers; one is at $\theta_1 = -A$ and the other is at $\theta_1 = A$. Therefore, sets \mathcal{L}_1 and \mathcal{L}_0 in Proposition 2 are defined as $\mathcal{L}_1 = \{-A, A\}$ and $\mathcal{L}_0 = \{0\}$, respectively. Hence, the conditions in Proposition 2 must hold for two groups: $f^{(2)}, f^{(2)}_A, g^{(2)}_0$ and $f^{(2)}, f^{(2)}_{-A}, g^{(2)}_0$. From Table 4, it is noted that

³Because scalar observations are considered, the signs of $f^{(2)}$, $f^{(2)}_{\theta_1}$, and $g^{(2)}_{\theta_0}$ in (22)-(24) do not depend on \mathbf{z} ; hence, $\mathbf{z} = 1$ is used for Table 4.

Table 3: Optimal additive noise p.d.f.s, in the form of $p_N(n) = \lambda_1 \,\delta(n-n_1) + \lambda_2 \,\delta(n-n_2) + (1-\lambda_1-\lambda_2) \,\delta(n-n_3)$, for various values of σ for $\beta = 0.78$, $\alpha = 0.45$, A = 1 and $\rho = 0.8$.

σ	λ_1	λ_2	n_1	n_2	<i>n</i> ₃
0	0.4510	0.12	0.2209	-0.2763	0.4344
0.05	0.5888	0.2912	0.2955	0.2848	-0.2895
0.15	0.7734	0.2266	0.2547	-0.2547	_
0.35	1	_	0.0608	_	_
0.45	1	_	0.0238	_	_

Table 4: Numerical values of the auxiliary functions defined for Proposition 2.

σ	$f^{(1)}$	$f_A^{(1)}$	$f_{-A}^{(1)}$	$g_0^{(1)}$	$f^{(2)}$	$f_{A}^{(2)}$	$f_{-A}^{(2)}$	$g_0^{(2)}$
0.05	0.1614	0.2694	-0.2705	0.0011	10.8	10.8	10.8	-21.6
0.10	0.3627	0.6046	-0.6052	6.049×10^{-4}	6.0489	6.0489	6.049	-12.1
0.15	0.3225	0.5376	-0.5378	2.25×10^{-4}	2.25	2.25	2.25	-4.5
0.20	0.2905	0.4841	-0.4842	5.502×10^{-5}	0.5507	0.5507	0.5507	-1.1
0.25	0.2856	0.4759	-0.4759	-2.758×10^{-5}	-0.2669	-0.2669	-0.2669	0.5515
0.30	0.2683	0.4772	-0.4771	-5.764×10^{-5}	-0.5395	-0.5395	-0.5395	1.153

 $f^{(2)}$, $f^{(2)}_A$ and $f^{(2)}_{-A}$ are always positive whereas $g^{(2)}_0$ is always negative for the given values of σ . For this reason, the third condition in Proposition 2 is satisfied for both groups for those values of σ , implying that the detector is improvable as a result of the proposition, which is also verified from Figures 1–3.

Finally, the conditions in Proposition 5 are checked in the following. We consider the Gaussian mixture noise in (43) with $\sigma = 0.05$, and calculate the values of $f^{(1)}$, $f^{(1)}_A$, $f^{(1)}_{-A}$, $g^{(1)}_0$, $f^{(2)}_A$, $f^{(2)}_A$, $f^{(2)}_{-A}$, and $g^{(2)}_0$. These values are tabulated in Table 4. From the signs of the first derivatives it is straightforward to construct the following sets:

• $S_{\beta}^{z} = \emptyset$, $S_{\beta}^{n} = -A$, $S_{\beta}^{p} = \{f^{(1)}, A\}$

•
$$\mathcal{S}^{z}_{\alpha} = \emptyset, \, \mathcal{S}^{n}_{\alpha} = \emptyset, \, \mathcal{S}^{p}_{\alpha} = \{0\}$$

Now the conditions in Proposition 5 are checked.

- 1. Since both S_{β}^{z} and S_{α}^{z} are empty sets, the first condition is automatically satisfied.
- 2. The first bullet of the second condition is not satisfied. Since $|S_{\beta}^{n}| + |S_{\alpha}^{n}| = 1$ is an odd number, we

have to check the condition in the second bullet, which reduces, for this example, to the following:

$$\min\{F_2(-A)F_1(A)G_1(0)f^{(1)}, G_2(0)F_1(A)F_1(-A)f^{(1)}\} > \\ \max\{f^{(2)}F_1(A)F_1(-A)G_1(0), F_2(A)F_1(-A)G_1(0)f^{(1)}\}.$$
(49)

Due to the signs of the derivatives, it turns out that the two inputs of the min function on the left-hand side are positive whereas the two inputs of the max function on the right-hand side are negative so that the inequality is satisfied.

Hence, the detector is improvable as a result of Proposition 5. Moreover, when $\sigma = 0.10$, $\sigma = 0.15$, or $\sigma = 0.20$, the signs of the derivatives are the same as those in the case of $\sigma = 0.05$. Therefore, for all these cases the detector is improvable.

Now consider the case in which $\sigma = 0.25$. Again, the values of $f^{(1)}$, $f_A^{(1)}$, $f_{-A}^{(1)}$, $g_0^{(1)}$, $f^{(2)}$, $f_A^{(2)}$, $f_{-A}^{(2)}$, and $g_0^{(2)}$ are tabulated in Table 4. In this scenario, the sets are obtained as follows:

- $S_{\beta}^{z} = \emptyset, S_{\beta}^{n} = -A, S_{\beta}^{p} = \{f^{(1)}, A\}$
- $S^z_{\alpha} = \emptyset, S^n_{\alpha} = \{0\}, S^p_{\alpha} = \emptyset$

Then, the conditions in Proposition 5 are checked as follows:

- 1. Since both S_{β}^{z} and S_{α}^{z} are empty sets, the first condition is satisfied.
- 2. The first bullet of the second condition is not satisfied. Since $|S_{\beta}^{n}| + |S_{\alpha}^{n}| = 2$ is an even number, we have to check the condition in the third bullet, which, reduces, for this example, to the following:

$$\min\{F_2(A)F_1(-A)G_1(0)f^{(1)}, G_2(0)F_1(A)F_1(-A)f^{(1)}, f^{(2)}F_1(A)F_1(-A)G_1(0)\} > \max\{F_2(-A)F_1(A)G_1(0)f^{(1)}\}$$
(50)

For this case it turns out that all three inputs of the min function on the left-hand side are positive and the single input to the max function on the right-hand side is negative so that the inequality is not satisfied. Hence, the improvability conditions in Proposition 5 are not satisfied for this scenario. Similar calculations show that the same holds for $\sigma = 0.30$ as well.

5. Concluding Remarks

Noise enhanced hypothesis-testing has been studied in the restricted NP framework. A problem formulation has been presented for the p.d.f. of optimal additive noise. Generic improvability and nonimprovability conditions have been derived to determine if additive noise can provide performance improvements over cases in which no additive noise is employed. Also, when the number of possible parameter values is finite, it has been stated that the optimal additive noise can be represented by a discrete random variable with a certain number of point masses. In addition, more specific improvability conditions have been derived for this scenario. Finally, the theoretical results have been investigated over a numerical example and improvements via additive noise have been illustrated.

Appendix A. Appendices

Appendix A.1. Proof of Proposition 2

For the improvability of a detector in the restricted NP framework, there must exist a noise p.d.f. $p_{\mathbf{N}}(\mathbf{n})$ that satisfies $\mathbb{E}\{F(\mathbf{N})\} > F(\mathbf{0})$, $\min_{\theta \in \Lambda_1} \mathbb{E}\{F_{\theta}(\mathbf{N})\} \ge \beta$, and $\max_{\theta \in \Lambda_0} \mathbb{E}\{G_{\theta}(\mathbf{N})\} \le \alpha$, which can be expressed as $\int_{\mathbb{R}^K} p_{\mathbf{N}}(\mathbf{n}) F(\mathbf{n}) d\mathbf{n} > F(\mathbf{0})$, $\int_{\mathbb{R}^K} p_{\mathbf{N}}(\mathbf{n}) F_{\theta}(\mathbf{n}) d\mathbf{n} \ge \beta$, $\forall \theta \in \Lambda_1$, and $\int_{\mathbb{R}^K} p_{\mathbf{N}}(\mathbf{n}) G_{\theta}(\mathbf{n}) d\mathbf{n} \le \alpha$, $\forall \theta \in \Lambda_0$. Employing a similar approach to that in the proof of Theorem 2 in [7], we consider a noise p.d.f. with L infinitesimal noise components, $p_{\mathbf{N}}(\mathbf{n}) = \sum_{j=1}^L \lambda_j \delta(\mathbf{n} - \boldsymbol{\epsilon}_j)$. Then, the conditions above become

$$\sum_{j=1}^{L} \lambda_j F(\boldsymbol{\epsilon}_j) > F(\mathbf{0}) , \quad \sum_{j=1}^{L} \lambda_j F_{\theta}(\boldsymbol{\epsilon}_j) \ge \beta , \ \forall \theta \in \Lambda_1 , \quad \sum_{j=1}^{L} \lambda_j G_{\theta}(\boldsymbol{\epsilon}_j) \le \alpha , \ \forall \theta \in \Lambda_0 .$$
 (A.1)

As $\boldsymbol{\epsilon}_{j}$'s are infinitesimally small, $F(\boldsymbol{\epsilon}_{j})$, $F_{\theta}(\boldsymbol{\epsilon}_{j})$, and $G_{\theta}(\boldsymbol{\epsilon}_{j})$ can be approximated via the Taylor series expansion as $F(\mathbf{0}) + \boldsymbol{\epsilon}_{j}^{T}\mathbf{f} + 0.5 \boldsymbol{\epsilon}_{j}^{T}\mathbf{H}\boldsymbol{\epsilon}_{\theta} + 0.5 \boldsymbol{\epsilon}_{j}^{T}\mathbf{H}_{\theta}^{\mathbf{f}}\boldsymbol{\epsilon}_{j}$, and $G_{\theta}(\mathbf{0}) + \boldsymbol{\epsilon}_{j}^{T}\mathbf{g}_{\theta} + 0.5 \boldsymbol{\epsilon}_{j}^{T}\mathbf{H}_{\theta}^{\mathbf{g}}\boldsymbol{\epsilon}_{j}$, respectively, where \mathbf{H} ($\mathbf{H}_{\theta}^{\mathbf{f}}, \mathbf{H}_{\theta}^{\mathbf{g}}$) and \mathbf{f} ($\mathbf{f}_{\theta}, \mathbf{g}_{\theta}$) are the Hessian and the gradient of $F(\mathbf{x})$ ($F_{\theta}(\mathbf{x}), G_{\theta}(\mathbf{x})$) at $\mathbf{x} = \mathbf{0}$, respectively.

tively. Hence, (A.1) leads to

$$\sum_{j=1}^{L} \lambda_{j} \boldsymbol{\epsilon}_{j}^{T} \mathbf{H} \boldsymbol{\epsilon}_{j} + 2 \sum_{j=1}^{L} \lambda_{j} \boldsymbol{\epsilon}_{j}^{T} \mathbf{f} > 0,$$

$$\sum_{j=1}^{L} \lambda_{j} \boldsymbol{\epsilon}_{j}^{T} \mathbf{H}_{\theta}^{\mathbf{f}} \boldsymbol{\epsilon}_{j} + 2 \sum_{j=1}^{L} \lambda_{j} \boldsymbol{\epsilon}_{j}^{T} \mathbf{f}_{\theta} \ge 2 \left(\beta - F_{\theta}(\mathbf{0})\right), \quad \forall \theta \in \Lambda_{1},$$

$$\sum_{j=1}^{L} \lambda_{j} \boldsymbol{\epsilon}_{j}^{T} \mathbf{H}_{\theta}^{\mathbf{g}} \boldsymbol{\epsilon}_{j} + 2 \sum_{j=1}^{L} \lambda_{j} \boldsymbol{\epsilon}_{j}^{T} \mathbf{g}_{\theta} \le 2 \left(\alpha - G_{\theta}(\mathbf{0})\right), \quad \forall \theta \in \Lambda_{0}.$$
(A.2)

Express ϵ_j as $\epsilon_j = \rho_j \mathbf{z}$ for j = 1, 2, ..., L, where ρ_j for j = 1, 2, ..., L are infinitesimal real numbers, and \mathbf{z} is a *K*-dimensional real vector. Then, based on the definitions in (19)-(24), the conditions in (A.2) can be simplified to the following:

$$\left(f^{(2)}(\mathbf{x}, \mathbf{z}) + c f^{(1)}(\mathbf{x}, \mathbf{z})\right)\Big|_{\mathbf{x}=\mathbf{0}} > 0$$
, (A.3)

$$\left(f_{\theta}^{(2)}(\mathbf{x}, \mathbf{z}) + c f_{\theta}^{(1)}(\mathbf{x}, \mathbf{z})\right)\Big|_{\mathbf{x}=\mathbf{0}} > \frac{2(\beta - F_{\theta}(\mathbf{0}))}{\sum_{j=1}^{L} \lambda_{j} \rho_{j}^{2}}, \ \forall \theta \in \Lambda_{1},$$
(A.4)

$$\left(g_{\theta}^{(2)}(\mathbf{x}, \mathbf{z}) + c g_{\theta}^{(1)}(\mathbf{x}, \mathbf{z})\right)\Big|_{\mathbf{x}=\mathbf{0}} < \frac{2\left(\alpha - G_{\theta}(\mathbf{0})\right)}{\sum_{j=1}^{L} \lambda_{j} \rho_{j}^{2}} , \ \forall \theta \in \Lambda_{0} ,$$
(A.5)

where $c \triangleq 2\sum_{j=1}^{L} \lambda_j \rho_j / \sum_{j=1}^{L} \lambda_j \rho_j^2$. Because $\beta = F_{\theta}(\mathbf{0})$ for $\theta \in \mathcal{L}_1$ ($\alpha = G_{\theta}(\mathbf{0})$ for $\theta \in \mathcal{L}_0$) and $\beta < \min_{\theta \in \Lambda_1 \setminus \mathcal{L}_1} F_{\theta}(\mathbf{0})$ ($\alpha > \max_{\theta \in \Lambda_0 \setminus \mathcal{L}_0} G_{\theta}(\mathbf{0})$), the right-hand-side of (A.4) ((A.5)) goes to minus infinity for $\{\theta \in \Lambda_1 \mid \theta \notin \mathcal{L}_1\}$ (plus infinity for $\{\theta \in \Lambda_0 \mid \theta \notin \mathcal{L}_0\}$). Hence, we should consider only the $\theta \in \mathcal{L}_1$ case for $\theta \in \Lambda_1$ and the $\theta \in \mathcal{L}_0$ case for $\theta \in \Lambda_0$. Thus, (A.3), (A.4), and (A.5) can be expressed as

$$\left(f^{(2)}(\mathbf{x}, \mathbf{z}) + c f^{(1)}(\mathbf{x}, \mathbf{z})\right)\Big|_{\mathbf{x}=\mathbf{0}} > 0$$
(A.6)

$$\left(f_{\theta_1}^{(2)}(\mathbf{x}, \mathbf{z}) + c f_{\theta_1}^{(1)}(\mathbf{x}, \mathbf{z})\right)\Big|_{\mathbf{x}=\mathbf{0}} > 0$$
(A.7)

$$\left(g_{\theta_0}^{(2)}(\mathbf{x}, \mathbf{z}) + c g_{\theta_0}^{(1)}(\mathbf{x}, \mathbf{z})\right)\Big|_{\mathbf{x}=\mathbf{0}} < 0.$$
(A.8)

Note that *c* can take any real value by definition via the selection of appropriate λ_i and infinitesimal ρ_i values for i = 1, 2, ..., L. Then, based on (A.6)-(A.8), the following conclusions are made for the three bullets in the proposition:

- If the conditions in the first bullet of Proposition 2 are satisfied, *c* can be set to a sufficiently large positive number to satisfy the inequalities in (A.6)-(A.8).
- If the conditions in the second bullet of Proposition 2 are satisfied, *c* can be set to a sufficiently large negative number to satisfy the inequalities in (A.6)-(A.8).
- If the conditions in the first bullet of Proposition 2 are satisfied, *c* can set to zero to satisfy the inequalities in (A.6)-(A.8).

Appendix A.2. Proof of Proposition 3

As J(t) in (25) and H(t) in (26) are second-order continuously differentiable around $t = \tilde{\alpha}$, one can find $\epsilon > 0$, \mathbf{n}_1 , and \mathbf{n}_2 such that $\max_{\theta \in \Lambda_0} G_{\theta}(\mathbf{n}_1) = \tilde{\alpha} + \epsilon$ and $\max_{\theta \in \Lambda_0} G_{\theta}(\mathbf{n}_2) = \tilde{\alpha} - \epsilon$ [7]. Then, in the following, it is proved that an additive noise component with $p_N(\mathbf{n}) = 0.5 \,\delta(\mathbf{x} - \mathbf{n}_1) + 0.5 \,\delta(\mathbf{x} - \mathbf{n}_2)$ improves the detector performance according to the restricted NP criterion (i.e., under the worst-case detection and false alarm constraints). First, under the condition of $H''(\tilde{\alpha}) \ge 0$, the minimum value of the detection probability and the maximum value of the false alarm probability in the presence of additive noise are shown not to remain below β and exceed α , respectively:

$$\min_{\theta \in \Lambda_1} \mathbb{E}\{F_{\theta}(\mathbf{N})\} \ge \mathbb{E}\left\{\min_{\theta \in \Lambda_1} F_{\theta}(\mathbf{N})\right\} \ge 0.5H(\tilde{\alpha} + \epsilon) + 0.5H(\tilde{\alpha} - \epsilon) \ge H(\tilde{\alpha}) = \tilde{\beta} \ge \beta$$
(A.9)

$$\max_{\theta \in \Lambda_0} \mathbb{E}\{G_{\theta}(\mathbf{N})\} \le \mathbb{E}\left\{\max_{\theta \in \Lambda_0} G_{\theta}(\mathbf{N})\right\} = 0.5(\tilde{\alpha} + \epsilon) + 0.5(\tilde{\alpha} - \epsilon) = \tilde{\alpha} \le \alpha .$$
(A.10)

In addition, due to the assumptions in the proposition, J(t) is convex in an interval around $t = \tilde{\alpha}$. As $E\{F(\mathbf{N})\}$ can achieve the value of $0.5 J(\tilde{\alpha} + \epsilon) + 0.5 J(\tilde{\alpha} - \epsilon)$, which is always larger than $J(\tilde{\alpha})$ due to convexity, it is concluded that $E\{F(\mathbf{N})\} > J(\tilde{\alpha})$. Since $J(\tilde{\alpha}) \ge F(\mathbf{0})$ by definition of J(t) in (25), $E\{F(\mathbf{N})\} > F(\mathbf{0})$ is satisfied. Therefore, the detector is improvable. \Box

Appendix A.3. Proof of Proposition 5

A similar approach to the proof of Theorem 2 in [7] can be employed. According to Proposition 4, the optimal additive noise has a discrete probability distribution with at most M + N point masses. Then, a

detector is improvable if there exists a noise p.d.f. $p_{\mathbf{N}}(\mathbf{n}) = \sum_{l=1}^{M+N} \lambda_l \,\delta(\mathbf{n} - \mathbf{n}_l)$ that satisfies $\mathbb{E}\{F(\mathbf{N})\} > F(\mathbf{0})$, $\min_{i \in \{1, 2, \dots, N\}} \mathbb{E}\{F_{\theta_{1i}}(\mathbf{N})\} \ge \beta$, and $\max_{i \in \{1, 2, \dots, M\}} \mathbb{E}\{G_{\theta_{0i}}(\mathbf{N})\} \le \alpha$, which can be stated as

$$\sum_{l=1}^{M} \lambda_l F(\mathbf{n}_l) > F(\mathbf{0})$$

$$\min_{i \in \{1, 2, \dots, M\}} \sum_{l=1}^{M+N} \lambda_l F_{\theta_{1i}}(\mathbf{n}_l) \ge \beta$$

$$\max_{i \in \{1, 2, \dots, M\}} \sum_{l=1}^{M+N} \lambda_l G_{\theta_{0i}}(\mathbf{n}_l) \le \alpha .$$
(A.11)

Similarly to the approach in the proof of Proposition 2 in Appendix A.1, consider the improvability conditions in (A.11) for infinitesimal noise components, $\mathbf{n}_l = \boldsymbol{\epsilon}_l = \rho_l \mathbf{z}$ for l = 1, 2, ..., M + N, where ρ_l 's are infinitesimal real numbers, and \mathbf{z} is a *K*-dimensional real vector. Then, based on similar manipulations to those in Appendix A.1, the following conditions are obtained:

$$\left(f^{(2)}(\mathbf{x}, \mathbf{z}) + c f^{(1)}(\mathbf{x}, \mathbf{z}) \right) \Big|_{\mathbf{x}=\mathbf{0}} > 0$$
(A.12)

$$\left(f_{i}^{(2)}(\mathbf{x}, \mathbf{z}) + c f_{i}^{(1)}(\mathbf{x}, \mathbf{z})\right)\Big|_{\mathbf{x}=\mathbf{0}} > \frac{2\left(\beta - F_{\theta_{1i}}(\mathbf{0})\right)}{\sum_{j=1}^{M} \lambda_{j} \rho_{j}^{2}}, \quad i = 1, 2, \dots, N$$
(A.13)

$$\left(g_{i}^{(2)}(\mathbf{x}, \mathbf{z}) + c g_{i}^{(1)}(\mathbf{x}, \mathbf{z})\right)\Big|_{\mathbf{x}=\mathbf{0}} < \frac{2\left(\alpha - G_{\theta_{0i}}(\mathbf{0})\right)}{\sum_{j=1}^{M} \lambda_{j} \rho_{j}^{2}}, \quad i = 1, 2, \dots, M$$
(A.14)

where $c \triangleq 2 \sum_{j=1}^{M} \lambda_j \rho_j / \sum_{j=1}^{M} \lambda_j \rho_j^2$.

Because $F_{\theta_{1i}}(0) > \beta$, $\forall i \in \bar{S}_{\beta}$ and $G_{\theta_{0i}}(0) < \alpha$, $\forall i \in \bar{S}_{\alpha}$, the right-hand-side of (A.13) and (A.14) becomes minus infinity for $i \in \bar{S}_{\beta}$ and plus infinity for $i \in \bar{S}_{\alpha}$, respectively. Therefore, it is sufficient to consider $i \in S_{\beta}$ and $i \in S_{\alpha}$ only. Hence, (A.12)-(A.14) can be expressed as

$$\left(f^{(2)}(\mathbf{x}, \mathbf{z}) + c f^{(1)}(\mathbf{x}, \mathbf{z})\right)\Big|_{\mathbf{x}=\mathbf{0}} > 0$$
 (A.15)

$$\left(f_i^{(2)}(\mathbf{x}, \mathbf{z}) + c f_i^{(1)}(\mathbf{x}, \mathbf{z})\right)\Big|_{\mathbf{x}=\mathbf{0}} > 0, \ \forall i \in \mathcal{S}_\beta$$
(A.16)

$$\left(g_i^{(2)}(\mathbf{x}, \mathbf{z}) + c g_i^{(1)}(\mathbf{x}, \mathbf{z})\right)\Big|_{\mathbf{x}=\mathbf{0}} < 0, \ \forall i \in \mathcal{S}_{\alpha}.$$
(A.17)

From the definitions in (33) and (34), (A.15)-(A.17) can be written as

$$\left(\mathcal{F}_{2}(j) + c \mathcal{F}_{1}(j)\right)\Big|_{\mathbf{x}=\mathbf{0}} > 0 \text{ for } j = 1, 2, \dots, |\mathcal{S}_{\beta}| + 1$$
 (A.18)

$$\left(\mathcal{G}_2(j) + c \,\mathcal{G}_1(j)\right)\Big|_{\mathbf{x}=\mathbf{0}} < 0 \text{ for } j = 1, 2, \dots, |\mathcal{S}_\alpha|.$$
(A.19)

It is again observed that *c* can take any real value by selecting appropriate λ_i and infinitesimal ρ_i values for i = 1, 2, ..., M + N. Therefore, from (35) and (38), it is concluded that for the conditions in (A.18) and (A.19) to hold,

$$\mathcal{F}_{2}(j)\Big|_{\mathbf{x}=\mathbf{0}} > 0 \ \forall j \in \mathcal{S}_{\beta}^{z} \text{ and } \mathcal{G}_{2}(j)\Big|_{\mathbf{x}=\mathbf{0}} < 0 \ \forall j \in \mathcal{S}_{\alpha}^{z}$$
(A.20)

must be satisfied, which is the first condition in the proposition.

In addition to (A.20), one of the following conditions must be satisfied for the improvability conditions in (A.18) and (A.19) to hold:

When any three of |Sⁿ_β|, |Sⁿ_β|, |Sⁿ_α|, and |S^p_α| are zero, as stated in the first part of the second condition in Proposition 5, all the second terms that are nonzero in (A.18) and (A.19) are either all non-negative or all non-positive and the corresponding signs of the inequalities are the same. Therefore, there always exists a *c* that satisfies the improvability conditions in (A.18) and (A.19) when the first condition in Proposition 5 (cf. (A.20)) is satisfied.

When $|S_{\beta}^{n}| + |S_{\alpha}^{p}| = 0$, as stated in the first part of the second condition in Proposition 5, assume that $|S_{\alpha}^{n}|$ is an odd number (this does not reduce the generality of the result in the proposition). Then,

(A.18) and (A.19) can be stated after some manipulations as

$$\mathcal{F}_2(j)\Big|_{\mathbf{x}=\mathbf{0}} > 0, \ \forall j \in \mathcal{S}_\beta^z \tag{A.21}$$

$$\mathcal{G}_2(j)\Big|_{\mathbf{x}=\mathbf{0}} < 0, \ \forall j \in \mathcal{S}^z_\alpha \tag{A.22}$$

$$\left(\mathcal{F}_{2}(j)\prod_{l\in\mathcal{S}_{\beta}^{p}\cup\mathcal{S}_{\alpha}^{n}\setminus\{j\}}\left(\mathcal{F}_{1}(l)\mathcal{I}_{\mathcal{S}_{\beta}^{p}}(l)+\mathcal{G}_{1}(l)\mathcal{I}_{\mathcal{S}_{\alpha}^{n}}(l)\right)+c\prod_{l\in\mathcal{S}_{\beta}^{p}\cup\mathcal{S}_{\alpha}^{n}}\left(\mathcal{F}_{1}(l)\mathcal{I}_{\mathcal{S}_{\beta}^{p}}(l)+\mathcal{G}_{1}(l)\mathcal{I}_{\mathcal{S}_{\alpha}^{n}}(l)\right)\right)\Big|_{\mathbf{x}=\mathbf{0}}<0,\ \forall j\in\mathcal{S}_{\beta}^{p}$$
(A.23)

$$\left(\mathcal{G}_{2}(j)\prod_{l\in\mathcal{S}_{\beta}^{p}\cup\mathcal{S}_{\alpha}^{n}\setminus\{j\}}\left(\mathcal{F}_{1}(l)\mathcal{I}_{\mathcal{S}_{\beta}^{p}}(l)+\mathcal{G}_{1}(l)\mathcal{I}_{\mathcal{S}_{\alpha}^{n}}(l)\right)+c\prod_{l\in\mathcal{S}_{\beta}^{p}\cup\mathcal{S}_{\alpha}^{n}}\left(\mathcal{F}_{1}(l)\mathcal{I}_{\mathcal{S}_{\beta}^{p}}(l)+\mathcal{G}_{1}(l)\mathcal{I}_{\mathcal{S}_{\alpha}^{n}}(l)\right)\right)\Big|_{\mathbf{x}=\mathbf{0}}<0,\ \forall j\in\mathcal{S}_{\alpha}^{n}.$$
(A.24)

In obtaining (A.23) and (A.24), (A.18) and (A.19) are multiplied by $\prod_{l \in S_{\beta}^{p} \cup S_{\alpha}^{n} \setminus \{j\}} (\mathcal{F}_{1}(l) \mathcal{I}_{S_{\beta}^{p}}(l) + \mathcal{G}_{1}(l) \mathcal{I}_{S_{\alpha}^{n}}(l)),$ which is a positive (negative) quantity when $j \in S_{\alpha}^{n}$ ($j \in S_{\beta}^{p}$) since $|S_{\alpha}^{n}|$ is an odd number. The conditions in (A.21) and (A.22) are satisfied from the first condition in Proposition 5. Therefore, there always exists a *c* that satisfies the improvability conditions in (A.23) and (A.24) as the second terms and the sign of the inequalities in (A.23) and (A.24) are the same. When $|S_{\alpha}^{n}|$ is an even number, only the sign of the inequalities (A.23) and (A.24) change; hence, the same result is valid as well.

When $|S_{\beta}^{p}| + |S_{\alpha}^{n}| = 0$, as stated in the first part of the second condition in Proposition 5, via similar manipulations as in the previous paragraph, it can be proved that the detector is improvable with the first condition in Proposition 5.

• When $|S_{\beta}^{n}| + |S_{\alpha}^{n}|$ is an odd number, $|S_{\beta}^{n}| + |S_{\alpha}^{p}| > 0$, $|S_{\alpha}^{n}| + |S_{\beta}^{p}| > 0$, (A.18) and (A.19) can be written

$$\mathcal{F}_2(j)\Big|_{\mathbf{x}=\mathbf{0}} > 0, \ \forall j \in \mathcal{S}^z_\beta \tag{A.25}$$

$$\mathcal{G}_2(j)\Big|_{\mathbf{x}=\mathbf{0}} < 0, \ \forall j \in \mathcal{S}^z_\alpha \tag{A.26}$$

$$\begin{split} & \left(\left(\mathcal{F}_{2}(j)I_{S_{\beta}^{n}}(j) + \mathcal{G}_{2}(j)I_{S_{\alpha}^{p}}(j)\right) \prod_{l \in S_{\beta}^{n} \cup S_{\beta}^{p} \cup S_{\alpha}^{n} \cup S_{\alpha}^{p} \cup S_{\alpha}^{p} \cup S_{\alpha}^{p} \cup S_{\beta}^{p} \cup S_{\alpha}^{p} \cup S_{\beta}^{p}(l) + \mathcal{G}_{1}(l)I_{S_{\alpha}^{n} \cup S_{\beta}^{p}}(l) + \mathcal{G}_{1}(l)I_{S_{\alpha}^{n} \cup S_{\alpha}^{p}}(l) \right) \\ & + c \prod_{l \in S_{\beta}^{n} \cup S_{\beta}^{p} \cup S_{\alpha}^{n} \cup S_{\alpha}^{p}} \left(\mathcal{F}_{1}(l)I_{S_{\beta}^{n} \cup S_{\beta}^{p}}(l) + \mathcal{G}_{1}(l)I_{S_{\alpha}^{n} \cup S_{\alpha}^{p}}(l)\right) \right) \prod_{l \in S_{\beta}^{n} \cup S_{\beta}^{p} \cup S_{\alpha}^{n} \cup S_{\alpha}^{p} \cup S_{\alpha}^{p}} \left(\mathcal{F}_{1}(l)I_{S_{\alpha}^{n} \cup S_{\alpha}^{p}}(l) + \mathcal{G}_{1}(l)I_{S_{\alpha}^{n} \cup S_{\alpha}^{p}}(l) + \mathcal{G}_{1}(l)I_{S_{\alpha}^{n} \cup S_{\alpha}^{p}}(l) + \mathcal{G}_{1}(l)I_{S_{\alpha}^{n} \cup S_{\alpha}^{p}}(l) + \mathcal{G}_{1}(l)I_{S_{\alpha}^{n} \cup S_{\alpha}^{p}}(l) \right) \\ & + c \prod_{l \in S_{\beta}^{n} \cup S_{\beta}^{p} \cup S_{\alpha}^{n} \cup S_{\alpha}^{p}} \left(\mathcal{F}_{1}(l)I_{S_{\alpha}^{n} \cup S_{\alpha}^{p}}(l) + \mathcal{G}_{1}(l)I_{S_{\alpha}^{n} \cup S_{\alpha}^{p}}(l)\right) \right) \Big|_{\mathbf{x} = \mathbf{0}} < 0, \ \forall j \in S_{\beta}^{p} \cup S_{\alpha}^{n} . \tag{A.28}$$

In obtaining (A.27) and (A.28), (A.18) and (A.19) are multiplied by $\prod_{l \in S_{\beta}^{n} \cup S_{\beta}^{p} \cup S_{\alpha}^{n} \cup S_{\alpha}^{p} \setminus \{j\}} (\mathcal{F}_{1}(l) \mathcal{I}_{S_{\beta}^{n} \cup S_{\beta}^{p}}(l) + \mathcal{G}_{1}(l) \mathcal{I}_{S_{\alpha}^{n} \cup S_{\alpha}^{p}}(l))$, which is a positive (negative) quantity when $j \in S_{\beta}^{n} \cup S_{\alpha}^{n}$ ($j \in S_{\beta}^{p} \cup S_{\alpha}^{p}$) since $|S_{\beta}^{n}| + |S_{\alpha}^{n}|$ is an odd number. The conditions in (A.25) and (A.26) are satisfied from the first condition in the proposition. Also, under the condition in (41), there always exists a *c* that satisfies the improvability conditions in (A.27) and (A.28).

When |Sⁿ_β| + |Sⁿ_α| is an even number, |Sⁿ_β| + |S^p_α| > 0, and |Sⁿ_α| + |S^p_β| > 0 (A.18) and (A.19) can be expressed by four conditions similar to those in (A.25)-(A.28) with the only difference being that the signs of the inequalities in (A.27) and (A.28) are switched. In that scenario, the first and the second conditions are satisfied from the first condition in the proposition. In addition, under the condition in (42), there always exists a *c* that satisfies the third and the fourth conditions.

Appendix A.4. Proof of Corollary 1

Because $\min_{i \in \{1,2,\dots,N\}} F_{\theta_{1i}}(\mathbf{0}) > \beta$ and $\max_{i \in \{1,2,\dots,M\}} G_{\theta_{0i}}(\mathbf{0}) < \alpha$, the right-hand-side of (A.13) and (A.14) in the proof of Proposition 5 become minus infinity and plus infinity for any *i*, respectively. Then, it is sufficient

to consider the condition in (A.12) only; namely,

$$\left(f^{(2)}(\mathbf{x}, \mathbf{z}) + c f^{(1)}(\mathbf{x}, \mathbf{z})\right)\Big|_{\mathbf{x}=\mathbf{0}} > 0.$$
 (A.29)

This condition can be expressed as $\mathbf{z}^T \mathbf{H} \mathbf{z} + c \, \mathbf{z}^T \mathbf{f} > 0$ in terms of the gradient \mathbf{f} and the Hessian \mathbf{H} of $F(\mathbf{x})$ at $\mathbf{x} = \mathbf{0}$. As c can take any real value by definition as discussed before and as \mathbf{z} can be chosen arbitrarily small, the improvability condition is always satisfied if $\mathbf{f} \neq \mathbf{0}$. On the other hand, if $\mathbf{f} = \mathbf{0}$, the improvability condition becomes $\mathbf{z}^T \mathbf{H} \mathbf{z} > 0$. In that case, if $F(\mathbf{x})$ is not concave around $\mathbf{x} = \mathbf{0}$, \mathbf{H} is not negative semidefinite. Then, there exists \mathbf{z} such that $\mathbf{z}^T \mathbf{H} \mathbf{z} > 0$ is satisfied. Therefore, the detector is improvable.

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